

Dependence of MIMO System Performance on the Joint Properties of Angular Power

Terence Betlehem[†], Tharaka A. Lamahewa[†]
and Thushara D. Abhayapala^{†‡}

[†] Research School of Information Science and Engineering, Australian National University, Australia

[‡] Wireless Signal Processing Program, National ICT Australia Ltd.

Canberra, Australia

[terence.betlehem, tharaka.lamahewa, thushara.abhayapala]@anu.edu.au

Abstract—In this paper, we use a novel MIMO channel model to characterize the dependence of ergodic capacity and diversity order on the joint statistics of the angular power density. The scattering environment of a MIMO channel is characterized by a double directional angular power distribution, describing the power transferred in each direction from transmitter aperture to receiver aperture. Angular power, which is typically separable Kronecker-modelled, is here generalized to include joint distribution properties using well-known bivariate probability density functions. We show that the joint properties of the power density, namely the shape and the orientation of power distribution contours, have significant impact on capacity and diversity of non-line-of-sight (NLOS) channels.

I. INTRODUCTION

With recent development of practical MIMO systems, there is need to quantify the attainable MIMO performance over realistic channels. Many options for channel models are now available but either require ray-tracing, non-linear parametrizations, or restrict simulation to fixed array geometries at the transmitter and receiver. We present a model for arbitrary array geometries, and use it to characterize the effect of non-Kronecker channel properties on MIMO system performance, as a function of the angular power distribution.

A rich body of literature exists for the characterization of MIMO channels. For non-line-of-sight channels the channel gains are dominated by their second order statistics [1]. To model the second order channel statistics, authors commonly use geometric models based on parametrizations of the double directional power distribution [2]. However, many authors still use the oversimplified Kronecker model, where signal correlation at the transmitter and receiver are assumed independent. This model poorly estimates the capacity [3] and antenna correlations [4] in many cases.

In this paper, we study the joint power distribution parameters on the performance of a MIMO system. We use a novel model for non-line-of-sight channels that allows an antenna geometry-free study of the channel. We quantify the effect on ergodic capacity and diversity order of (i) the level of correlation between transmit and receive power spectra and (ii) the orientation of distribution contours in the transmit-angle/receive-angle plane.

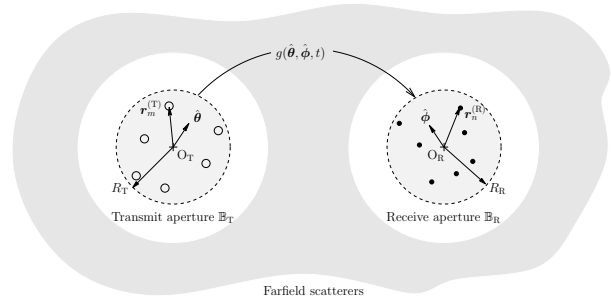


Fig. 1. The channel model. Scattering is modelled with the scattering gain $g(\theta, \phi, t)$. Each transmitter (\circ) is positioned at $\mathbf{r}_m^{(T)}$ within circular aperture of radii R_T and receiver (\cdot) at $\mathbf{r}_n^{(R)}$ within aperture of radius R_R .

II. CHANNEL MODEL

MIMO system performance is controlled by the correlation between the signals at the transmit and receive antennas. In this section, we model the channel correlation. We present a 2D model, which is reasonable for urban channel environments.

Consider transmission of data from n_T transmitter antennas over a flat fading channel to n_R receiver antennas. Let $\mathbf{r}_m^{(T)}$ be the m th transmitter antenna position with respect to a transmitter origin O_T and $\mathbf{r}_n^{(R)}$ be the n th receiver antenna position with respect to a receiver origin O_R (as shown in Fig. 1). Transmitter and receiver antennas lie within finite circular apertures of radius R_T and R_R respectively. Transmitters transmit the $n_T \times 1$ vector signal $\mathbf{x}(t)$ over the time-varying channel described by matrix $\mathbf{H}(t)$ to receive the $n_R \times 1$ vector signal $\mathbf{y}(t)$ at receiver antennas. Matrix $\mathbf{H}(t)$ houses the transfer functions $[\mathbf{H}]_{nm} \triangleq h_{nm}(t)$ from each transmitter antenna m to receiver antenna n . Defining the $n_R \times 1$ vector of additive white Gaussian noise $\mathbf{n}(t)$ at the receivers, we write:

$$\mathbf{y}(t) = \mathbf{H}(t)\mathbf{x}(t) + \mathbf{n}(t). \quad (1)$$

Assume that all scatterers lie in the far-field. The scattering environment causes transmitter signals to propagate in as plane waves with a different amplitude for each direction. Define the *scattering gain* $g(\theta, \phi, t)$ as the complex gain at time t of the signal propagating out from the transmitter origin in direction θ and arrives at the receiver origin in direction ϕ . We can then

write:

$$h_{nm}(t) = \int_{2\pi} \int_{2\pi} g(\theta, \phi, t) e^{ik(\mathbf{x}_m \cdot \hat{\theta} - \mathbf{y}_n \cdot \hat{\phi})} d\phi, \quad (2)$$

where $\hat{\theta}$ and $\hat{\phi}$ are unit vectors of polar coordinates $(1, \phi)$ and $(1, \theta)$ and k is the wave number. Assume slow fading so that the channel remains static over the symbol time. Then h_{nm} and g are not dependent on t over each symbol.

To investigate the statistics of non-line-of-sight scattering, $g(\theta, \phi)$ is assumed to be a zero mean Gaussian random variable at each angle pair (θ, ϕ) . Assume $g(\theta, \phi)$ is uncorrelated between different angles:

$$E_g\{g(\theta, \phi)g^*(\vartheta, \varphi)\} = \mathcal{P}(\theta, \phi)\delta_{\theta\vartheta}\delta_{\phi\varphi}, \quad (3)$$

where $\delta_{\phi\varphi}$ is the Kronecker delta function and $*$ is the complex conjugate and $\mathcal{P}(\theta, \phi) \triangleq E_g\{|g(\theta, \phi)|^2\}$ is the power density of scatterers, interpreted as the average energy transmitted in direction θ and received in direction ϕ .¹ The expectation is computed over all possible realizations of the scattering environment $g(\theta, \phi)$. Equation 3 is referred to as the wide-sense stationary uncorrelated scatterer (WSSUS) assumption [5].

Without loss of generality, we normalize the average energy of each channel coefficient to 1, which is equivalent to normalizing $\mathcal{P}(\theta, \phi)$:

$$\int_{2\pi} \int_{2\pi} \mathcal{P}(\theta, \phi) d\theta d\phi = 1.$$

Statistics of $g(\theta, \phi)$ and \mathbf{H} are dependent entirely upon $\mathcal{P}(\theta, \phi)$.

The univariate marginals of the angular power are understood as the power density of scatterers at the transmitter:

$$\mathcal{P}_T(\theta) \triangleq \int_{2\pi} \mathcal{P}(\theta, \phi) d\phi, \quad (4a)$$

and power density of scatterers at the receiver:

$$\mathcal{P}_R(\phi) \triangleq \int_{2\pi} \mathcal{P}(\theta, \phi) d\theta. \quad (4b)$$

Suitable functional forms for the univariate scatterer distributions have been well-studied [6]. These distributions characterize the properties of the separable Kronecker model, where it can be shown that $\mathcal{P}(\theta, \phi) = \mathcal{P}_T(\theta)\mathcal{P}_R(\phi)$ [4]. This paper relaxes the Kronecker assumption, by additionally describing power density with a jointness parameter.

The MIMO model is used to calculate the second order statistics $E_g\{h_{nm}h_{n'm'}^*\}$ of the entries of \mathbf{H} . These statistics can be collected in the channel correlation matrix $\mathbf{R}_H \triangleq E_g\{\vec{\mathbf{H}}\vec{\mathbf{H}}^\dagger\}$ where vectorize operator $\vec{\cdot}$ stacks the columns of a matrix and \cdot^\dagger is the Hermitian operator.

III. FOURIER ANALYSIS OF CHANNEL MATRIX

We now derive a compact expression for \mathbf{R}_H . By transforming functions into the Fourier domain, we decompose the channel matrices \mathbf{H} and \mathbf{R}_H into components depending separately on the antenna geometry and scattering environment.

¹Power density of scatterers is not to be confused with the probability density function of the scattering gain, which we model as complex Gaussian.

A. Fourier Decomposition

Perform the double Fourier expansion of $\mathcal{P}(\theta, \phi)$,

$$\mathcal{P}(\theta, \phi) = \frac{1}{(2\pi)^2} \sum_{\ell=-\infty}^{\infty} \sum_{\ell'=-\infty}^{\infty} \gamma_{\ell\ell'} e^{i(\ell\theta - \ell'\phi)}, \quad (5a)$$

$$\gamma_{\ell\ell'} = \int_{2\pi} \int_{2\pi} \mathcal{P}(\theta, \phi) e^{-i(\ell\theta - \ell'\phi)} d\theta d\phi, \quad (5b)$$

where $\gamma_{\ell\ell'}$ is the Fourier coefficient of $\mathcal{P}(\theta, \phi)$. Since $\mathcal{P}(\theta, \phi) \in \mathbb{R}$, we know that $\gamma_{\ell(-\ell')}^* = \gamma_{(-\ell)\ell'}$. Substituting (5a) into (4) we see the power density of scatterers around transmitter and receiver are a function of $\gamma_{\ell 0}$ and $\gamma_{0\ell'}$:

$$\mathcal{P}_T(\theta) = \frac{1}{2\pi} \sum_{\ell=-\infty}^{\infty} \gamma_{\ell 0} e^{i\ell\theta},$$

$$\mathcal{P}_R(\phi) = \frac{1}{2\pi} \sum_{\ell'=-\infty}^{\infty} \gamma_{0\ell'} e^{-i\ell'\phi}.$$

Similarly write $g(\theta, \phi)$ as the double Fourier expansion:

$$g(\theta, \phi) = \frac{1}{(2\pi)^2} \sum_{\ell=-\infty}^{\infty} \sum_{\ell'=-\infty}^{\infty} \beta_{\ell\ell'} e^{i(\ell\theta - \ell'\phi)}, \quad (7a)$$

$$\beta_{\ell\ell'} = \int_{2\pi} \int_{2\pi} g(\theta, \phi) e^{-i(\ell\theta - \ell'\phi)} d\theta d\phi, \quad (7b)$$

where $\beta_{\ell\ell'}$ is the Fourier coefficient of $g(\theta, \phi)$.

Due to the limited aperture sizes, (7a) can be truncated and from (2) each $h_{nm}(t)$ written as a finite sum of the $\beta_{\ell\ell'}$ coefficients. Drawing from [7], the channel matrix \mathbf{H} can be decomposed into three matrices:

$$\mathbf{H} = \mathbf{J}_R \boldsymbol{\beta} \mathbf{J}_T^\dagger, \quad (8)$$

where

$$\boldsymbol{\beta} = \begin{bmatrix} \beta_{(-N_R)(-N_T)} & \cdots & \beta_{(-N_R)N_T} \\ \beta_{(-N_R+1)(-N_T)} & \cdots & \beta_{(-N_R+1)N_T} \\ \vdots & \ddots & \vdots \\ \beta_{N_R(-N_T)} & \cdots & \beta_{N_R N_T} \end{bmatrix}, \quad (9)$$

is a matrix which completely characterizes the effect of the scattering environment, \mathbf{J}_T and \mathbf{J}_R are matrices describing the effect of the transmitter and receiver antenna geometries

$$\mathbf{J}_S = \begin{bmatrix} \mathcal{J}_{-N_S}(\mathbf{r}_1^{(S)}) & \cdots & \mathcal{J}_{N_S}(\mathbf{r}_1^{(S)}) \\ \vdots & \ddots & \vdots \\ \mathcal{J}_{-N_S}(\mathbf{r}_{n_S}^{(S)}) & \cdots & \mathcal{J}_{N_S}(\mathbf{r}_{n_S}^{(S)}) \end{bmatrix}, \quad (10)$$

where $S = T$ for \mathbf{J}_T and $S = R$ for \mathbf{J}_R , and if vector \mathbf{r} is defined in polar coordinates as (r, θ_r) , then

$$\mathcal{J}_\ell(\mathbf{r}) \triangleq (-i)^\ell J_\ell(kr) e^{i\ell\theta_r}, \quad (11)$$

$N_T = \lceil ekR_T/2 \rceil$ and $N_R = \lceil ekR_R/2 \rceil$ are the aperture dimensionalities of transmitter and receiver respectively and $J_m(\cdot)$ is the Bessel function of the first kind of order m .

We next model the scattering matrix $\boldsymbol{\beta}$ as a random matrix.

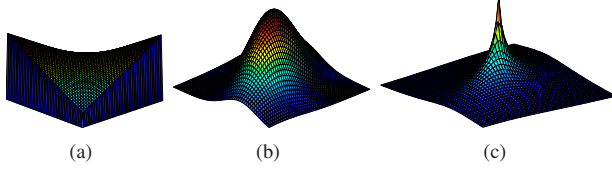


Fig. 2. Shapes of scatterer distributions used for joint distribution modelling: (a) Morgenstern (here $\rho_m = 0.7$), (b) bivariate Gaussian ($\theta_v = \pi/6$, $\sigma_1 = 2$, $\sigma_2 = 1$) and (c) bivariate Laplacian ($\theta_v = \pi/6$, $\sigma_1 = 4$, $\sigma_2 = 2$).

B. Statistics of Scattering Environment

Since the scattering gain is zero mean Gaussian, the statistics of the elements of $\vec{\beta}$ are also governed by its correlation matrix $\mathbf{R}_\beta \triangleq E_g\{\vec{\beta}\vec{\beta}^\dagger\}$. From (7b) each second order statistic $E_g\{\beta_{pp'}\beta_{qq'}^*\}$ can be calculated as:

$$E_g\{\beta_{pp'}\beta_{qq'}^*\} = \int_{2\pi} \int_{2\pi} \int_{2\pi} \int_{2\pi} e^{-i(p\theta - p'\phi)} e^{i(q\vartheta - q'\varphi)} \times E_g\{g(\theta, \phi)g^*(\vartheta, \varphi)\} d\theta d\phi d\vartheta d\varphi.$$

Applying the WSSUS property (3):

$$E_g\{\beta_{pp'}\beta_{qq'}^*\} = \int_{2\pi} \int_{2\pi} \mathcal{P}(\theta, \phi) e^{-i(p-q)\theta} e^{i(p'-q')\phi} d\theta d\phi.$$

We see by comparison with (5b) that $E_g\{\beta_{pp'}\beta_{qq'}^*\} = \gamma_{(p-q)(p'-q')}$ from which a block Toeplitz structure can be inferred for \mathbf{R}_β :

$$\mathbf{R}_\beta = \begin{bmatrix} \mathbf{\Gamma}_0 & \mathbf{\Gamma}_{-1} & \dots & \mathbf{\Gamma}_{-2N_T} \\ \mathbf{\Gamma}_1 & \mathbf{\Gamma}_0 & \dots & \mathbf{\Gamma}_{-2N_T+1} \\ \vdots & \vdots & \ddots & \vdots \\ \mathbf{\Gamma}_{2N_T} & \mathbf{\Gamma}_{2N_T-1} & \dots & \mathbf{\Gamma}_0 \end{bmatrix}, \quad (12)$$

where $\mathbf{\Gamma}_\ell$ is Toeplitz:

$$\mathbf{\Gamma}_\ell = \begin{bmatrix} \gamma_{0\ell} & \gamma_{(-1)\ell} & \dots & \gamma_{(-2N_R)\ell} \\ \gamma_{1\ell} & \gamma_{0\ell} & \dots & \gamma_{(-2N_R+1)\ell} \\ \vdots & \vdots & \ddots & \vdots \\ \gamma_{2N_R\ell} & \gamma_{(2N_R-1)\ell} & \dots & \gamma_{0\ell} \end{bmatrix}.$$

By virtue of the $\gamma_{\ell(-\ell')} = \gamma_{(-\ell)\ell'}^*$ property, $\mathbf{\Gamma}_{-\ell} = \mathbf{\Gamma}_\ell^\dagger$.

The statistics of \mathbf{H} are governed by covariance \mathbf{R}_β of β and the antenna geometry. From (8) and Kronecker relation $\overline{ABC} = (\mathbf{A}^T \otimes \mathbf{C})\overline{B}$ where \otimes is the Kronecker product and \cdot^T is the vector transpose operator [8], the channel correlation matrix \mathbf{R}_H is related to \mathbf{R}_β through

$$\mathbf{R}_H = (\mathbf{J}_T^* \otimes \mathbf{J}_R)\mathbf{R}_\beta(\mathbf{J}_T^* \otimes \mathbf{J}_R)^\dagger. \quad (13)$$

This equation separates antenna correlation into the antenna geometry term $\mathbf{J}_T^* \otimes \mathbf{J}_R$ and scattering environment correlation term \mathbf{R}_β . We can then investigate separately the contributions of either component to MIMO performance. Equation 13 describes both Kronecker and non-Kronecker channels.

IV. BI-ANGULAR POWER DISTRIBUTIONS

In this paper we consider the dependence on joint properties of the power distribution. Although various power distributions functional forms have been proposed for the univariate marginals at the transmitter and receiver, little has been suggested in terms of bivariate distributions for $\mathcal{P}(\theta, \phi)$. We propose natural extensions of the univariate marginals, the uniform limited, Gaussian and Laplacian distributions [6], to the bi-variate case. These distributions are parametrized by a mean angle-of-arrival θ_0 , mean angle-of-departure ϕ_0 , angular spread at the transmitter σ_t and receiver σ_r , and parameter ρ controlling joint properties. They are illustrated in Fig. 2.

A. Morgenstern Distributed Scatterers

In the case the energy leaves the transmitter aperture uniformly from $(\theta_0 - \Delta_t, \theta_0 + \Delta_t)$ to arrive at the receiver aperture uniformly from $(\theta_0 - \Delta_r, \theta_0 + \Delta_r)$, marginal constraints (4) can be satisfied by Morgenstern's family of distributions [9]:

$$f_M(\theta, \phi) = \frac{1}{4\Delta_t\Delta_r} - \frac{\rho_m(\theta - \theta_0)(\phi - \phi_0)}{4\Delta_t^2\Delta_r^2}.$$

Using (5b) it is straight-forward to derive:

$$\gamma_{\ell\ell'}^{(M)} = \begin{cases} e^{i\ell\theta_0}\text{sinc}(\ell\Delta_t), & \ell' = 0, \\ e^{-i\ell'\phi_0}\text{sinc}(\ell'\Delta_r), & \ell = 0, \\ e^{i(\ell\theta_0 - \ell'\theta_0)}\Gamma(\ell, \ell'), & \text{otherwise,} \end{cases}$$

where

$$\Gamma(\ell, \ell') = \text{sinc}(\ell\Delta_t)\text{sinc}(\ell'\Delta_r) - \frac{\rho}{\ell\Delta_t\ell'\Delta_r} \times [\cos(\ell\Delta_t) - \text{sinc}(\ell\Delta_t)][\cos(\ell'\Delta_r) - \text{sinc}(\ell'\Delta_r)].$$

Parameter $\rho \in [-1, 1]$ controls the shape of $f_M(\theta, \phi)$. In case $\rho = 0$, the distribution reduces to the separable product of two uniform limited distributions, the Kronecker case. Further in the isotropic case $\Delta_t = \pi$ and $\Delta_r = \pi$ the Morgenstern distribution reduces to the rich scattering distribution.

B. Gaussian Distributed Scatterers

The truncated bivariate Gaussian distribution is written:

$$f_G(\theta, \phi) = \Omega_G e^{-\frac{Q(\theta, \phi)}{2(1-\rho^2)}}, \quad |\theta - \theta_0|, |\phi - \phi_0| \leq \pi$$

where

$$Q(\theta, \phi) = \frac{(\theta - \theta_0)^2}{\sigma_t^2} - \frac{2\rho(\theta - \theta_0)(\phi - \phi_0)}{\sigma_t\sigma_r} + \frac{(\phi - \phi_0)^2}{\sigma_r^2},$$

ρ controls the level of coupling between the transmitter and receiver scatterers and Ω_G is a normalization constant. The Kronecker case is obtained by setting $\rho = 0$. For small angular spread ($\sigma_r, \sigma_t \ll \pi$), $\Omega_G = 1/2\pi\sqrt{1-\rho^2}\sigma_t\sigma_r$ and coefficients $\gamma_{\ell\ell'}$ are obtained from the Gaussian characteristic function [10]:

$$\gamma_{\ell\ell'}^{(G)} = e^{i(\ell\theta_0 - \ell'\phi_0)} e^{-\frac{1}{2}[\ell^2\sigma_t^2 + \ell'^2\sigma_r^2 - 2\sigma_t\sigma_r\rho\ell\ell']}.$$

The marginal distributions possess characteristic functions of form $\gamma_\ell^{(G)} = e^{i\ell\theta_0} e^{-\frac{1}{2}\ell^2\sigma_t^2}$, which correspond to the univariate

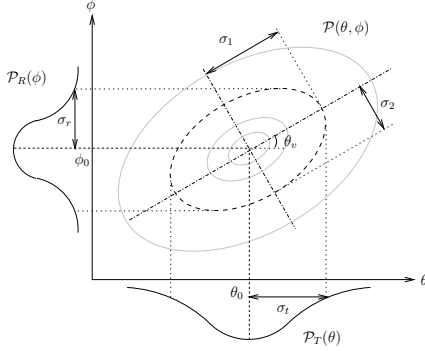


Fig. 3. Contour plot for a typical elliptical bivariate scatterer power distribution, showing distribution parameters: mean direction pair (θ_0, ϕ_0) and major and minor axes of ellipses σ_1 and σ_2 , respectively. Also shown are the marginal density of scatterers $P_T(\theta)$ and $P_R(\phi)$.

Gaussian distribution $f_G(\theta) = \frac{1}{\sqrt{2\pi}\sigma_t} e^{-(\theta-\theta_0)^2/2\sigma_t^2}$ [6]. The variances of the marginals, at the transmitter and receiver, are σ_t^2 and σ_r^2 respectively.

The distribution is an elliptically contoured function centered about (θ_0, ϕ_0) and parameterized in terms the ellipse orientation angle θ_v (see Fig. 3) and major and minor axis variances σ_1^2 and σ_2^2 respectively. The joint parameter ρ is a function of these geometric parameters [10]:

$$\rho = \frac{\sigma_1^2 - \sigma_2^2}{\sigma_t \sigma_r} \sin \theta_v \cos \theta_v,$$

where $\sigma_t^2 = \sigma_1^2 + (\sigma_2^2 - \sigma_1^2) \sin^2 \theta_v$ and $\sigma_r^2 = \sigma_2^2 + (\sigma_1^2 - \sigma_2^2) \sin^2 \theta_v$. Geometric parameters are illustrated in Fig. 3.

C. Laplacian Distributed Scatterers

Similarly to the Gaussian distribution, a truncated elliptical bivariate Laplacian distribution can also be written [11]:

$$f_L(\theta, \phi) = \Omega_L K_0 \left(\sqrt{\frac{Q(\theta, \phi)}{1 - \rho^2}} \right), |\theta - \theta_0|, |\phi - \phi_0| \leq \pi,$$

where $K_0(\cdot)$ is the modified Bessel function of the second kind of order zero and Ω_L is a normalization constant. For small angular spread, $\Omega_L = 1/\pi\sqrt{1 - \rho^2}\sigma_t\sigma_r$ and coefficients are obtained from the Laplacian characteristic function [11]:

$$\gamma_{\ell\ell'}^{(L)} = \frac{e^{i(\ell\theta_0 - \ell'\phi_0)}}{1 + \sigma_t^2 \ell^2 + \sigma_r^2 \ell'^2 - 2\rho\sigma_t\sigma_r\ell\ell'}.$$

The marginal distributions possess characteristic functions of form $\gamma_\ell^{(L)} = e^{i\ell\theta_0}/(1 + \sigma_t^2 \ell^2)$ which correspond to the well-known univariate Laplacian distribution $f_L(\theta) = \frac{1}{2\sigma_t} e^{-|\theta-\theta_0|/\sigma_t}$ with variance $2\sigma_t^2$.

V. INFORMATION THEORETIC PERFORMANCE MEASURES

We now describe the measures of capacity and diversity used for assessing MIMO system performance. The ergodic capacity of a MIMO channel with no channel state information

at the transmitter, is given by computing the mutual information with equally allotted transmit powers [12]:

$$C_{\text{erg}} = E_{\mathbf{H}} \left\{ \log_2 \left| \mathbf{I}_{n_T} + \frac{\text{SNR}}{n_R} \mathbf{H} \mathbf{H}^\dagger \right| \right\}, \quad (14)$$

where \mathbf{I}_{n_T} is the $n_T \times n_T$ identity matrix and SNR the average received signal-to-noise ratio.

Channel diversity is generally defined by the distribution of the $n_R n_T$ eigenvalues λ_i of the channel correlation matrix \mathbf{R}_H . These eigenvalues describe the average powers of the independent eigenvector channels of a MIMO channel. A useful single parameter diversity measure describing the eigenvalue spread is the diversity order. In [13], for an NLOS channel with Gaussian coefficients it was shown equal to:

$$\Psi(\mathbf{R}_H) = \frac{(\sum_{i=1}^{n_R n_T} \lambda_i)^2}{\sum_{i=1}^{n_R n_T} \lambda_i^2} = \left(\frac{\text{Tr}\{\mathbf{R}_H\}}{\|\mathbf{R}_H\|_F} \right)^2, \quad (15)$$

where $\|\cdot\|_F$ is the Frobenius norm and $\text{Tr}(\cdot)$ is the matrix trace. Here $1 < \Psi(\mathbf{R}_H) < n_T n_R$.

To obtain antenna geometry-independent performance results for our antenna apertures, we set $\mathbf{J}_T^\dagger \mathbf{J}_T = \mathbf{I}_{n_T}$ and $\mathbf{J}_R^\dagger \mathbf{J}_R = \mathbf{I}_{n_R}$. These relations are satisfied for $n_T \geq 2N_T + 1$ transmitter and $n_R \geq 2N_R + 1$ receiver antennas that are separated far enough apart so that the rich scattering ($\mathbf{R}_H = \mathbf{I}$) antenna correlation is negligible.

Using (8), we see immediately that:

$$\mathbf{H} \mathbf{H}^\dagger = \mathbf{J}_R \boldsymbol{\beta} \mathbf{J}_T^\dagger \mathbf{J}_T \boldsymbol{\beta}^\dagger \mathbf{J}_R^\dagger = (\mathbf{J}_R \boldsymbol{\beta}) (\boldsymbol{\beta}^\dagger \mathbf{J}_R^\dagger).$$

For fixed antennas, computing expectation (14) over channel realizations \mathbf{H} is equivalent to computing it over $g(\theta, \phi)$ or $\boldsymbol{\beta}$. Using the matrix identity $|\mathbf{I} + \mathbf{A}\mathbf{B}| = |\mathbf{I} + \mathbf{B}\mathbf{A}|$:

$$\begin{aligned} C_{\text{erg}} &= E_{\boldsymbol{\beta}} \left\{ \log_2 \left| \mathbf{I}_{n_T} + \frac{\text{SNR}}{n_R} \boldsymbol{\beta}^\dagger \mathbf{J}_R^\dagger \mathbf{J}_R \boldsymbol{\beta} \right| \right\} \\ &= E_{\boldsymbol{\beta}} \left\{ \log_2 \left| \mathbf{I}_{n_T} + \frac{\text{SNR}}{n_R} \boldsymbol{\beta}^\dagger \boldsymbol{\beta} \right| \right\}. \end{aligned} \quad (16)$$

Similarly for the diversity, using Kronecker product property $(\mathbf{A} \otimes \mathbf{B})(\mathbf{C} \otimes \mathbf{D}) = \mathbf{A}\mathbf{C} \otimes \mathbf{B}\mathbf{D}$:

$$(\mathbf{J}_T^* \otimes \mathbf{J}_R)^\dagger (\mathbf{J}_T^* \otimes \mathbf{J}_R) = \mathbf{J}_T^T \mathbf{J}_T^* \otimes \mathbf{J}_R^\dagger \mathbf{J}_R = \mathbf{I}_{n_T n_R}.$$

Hence from (13), \mathbf{R}_H and \mathbf{R}_β are similar matrices and possess the same trace and Frobenius norm. As a result, diversity order is given by:

$$\Psi(\mathbf{R}_H) = \left(\frac{\text{Tr}\{\mathbf{R}_\beta\}}{\|\mathbf{R}_\beta\|_F} \right)^2. \quad (17)$$

From (16) and (17), both capacity and diversity are hence characterized by the scattering matrix $\boldsymbol{\beta}$.

VI. SIMULATION OF JOINTLY CORRELATED CHANNELS

In the following examples, we use the above metrics to investigate the joint properties of the MIMO channel. Here antennas are packed into apertures of radius $R_T = R_R = \lambda$.

Ergodic capacity C_{erg} was calculated by ensemble averaging over 1000 realizations of the scatter matrix $\boldsymbol{\beta}$. SNR was

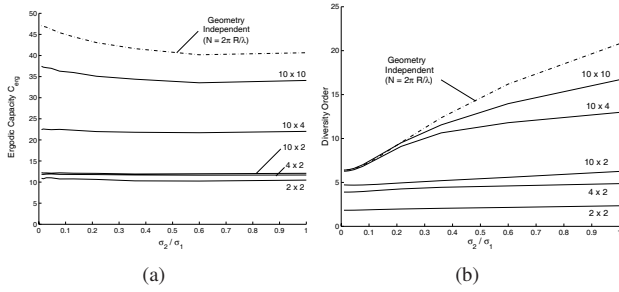


Fig. 4. (a) Ergodic capacity and (b) diversity order for various MIMO antenna schemes in a scattering environment with bivariate Gaussian angular power. Here $\theta_v = 45^\circ$, $\theta_t = \theta_r = 0.5$ while the major-minor axis ratio α is varied. Shown are geometry independent measures (dotted lines).

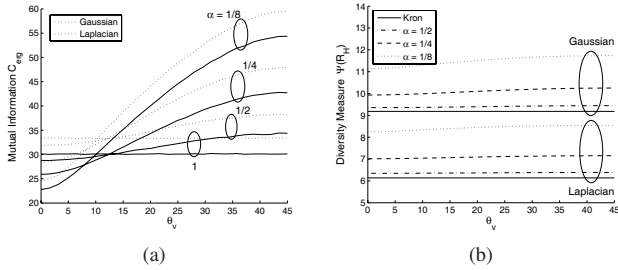


Fig. 5. (a) Ergodic capacity and (b) diversity order for Gaussian (Laplacian) distribution when varying the orientation angle θ_v , for constant ellipse area $\pi\sigma_1\sigma_2 = 0.1\pi$ (0.2π) but different major-minor axis ratio α .

set to 20dB. The set of channel scatter matrices $\{\beta_n\}$ was simulated from the coefficient correlation matrix \mathbf{R}_β by: (i) generating a series of $(2N_T + 1)(2N_R + 1)$ -length vectors of IID zero mean unit variance circularly symmetric Gaussian random variables $\{\vec{\mathbf{G}}_i\}$, and (ii) calculating realizations of the scatter matrix using $\vec{\beta}_i = \mathbf{R}_\beta^{1/2} \vec{\mathbf{G}}_i$.

A. Jointness of Power Distributions

We first study the effect of keeping the scatterer densities around the transmitter and receiver constant, while varying the joint parameter through the elliptical distribution parameters: ellipse major-axis-to-minor-axis ratio $\alpha \triangleq \sigma_2/\sigma_1$ and elliptical contour orientation θ_v .

Anticipating largest joint effects occur for ellipses oriented at $\theta_v = 45^\circ$, we plot the measures against α in Fig. 4. Here maximum distribution jointness ($\alpha = 0$) offers a modest 15% boost to capacity but a large drop in diversity. Though plotted for Gaussian scatterers with $\theta_t = \theta_r = 0.5$, same-shaped curves result for other angular spreads and Laplacian scatterers.

Fig. 4 also compares the geometry independent performance measures with those calculated directly from (14) and (15), for antenna configured into uniform circular arrays of radius λ . The geometry independent measures are reasonably close to the 10×10 results.

B. Orientation of the Distribution Contours

We also study the effect of orientation of the distribution contours, holding the contour area $\pi\sigma_1\sigma_2$ of the power dis-

tribution constant. In Fig. 5, we see capacity depends significantly on contour orientation, while diversity order is relatively fixed. Further simulations show that diversity order is even less strongly dependent upon orientation for smaller aperture sizes. Diversity order shows much stronger dependence upon contour area (which depends linearly on α) than contour orientation. However, contour orientation is shown important for modeling MIMO channel capacity.

VII. CONCLUSION

In this paper, we explore the impact of the joint properties of the angular power distribution on MIMO system performance. We present a novel NLOS modelling approach to generate an antenna geometry-free analysis of the MIMO channel. In doing so, we show that the joint properties can significantly effect diversity order and ergodic capacity. Capacity is strongly dependent upon distribution contour orientation while the diversity order proportional to the area of distribution contours.

ACKNOWLEDGMENT

This work was partially funded by Australian Research Council Discovery Grant number DP0343804.

REFERENCES

- [1] K. Yu, M. Bengtsson, B. Ottersten, D. McNamara, P. Karlsson, and M. Beach, "Second order statistics of NLOS indoor MIMO channels based on 5.2 GHz measurements," in *IEEE Global Communications Conf.*, Nov. 2001.
- [2] Q. H. Spencer, B. D. Jeffs, M. A. Jensen, and A. L. Swindlehurst, "Modeling the statistical time and angle of arrival characteristics of an indoor multipath channel," *IEEE J. Select. Areas Commun.*, vol. 18, no. 3, pp. 347 – 360, 2000.
- [3] H. Ozcelik, N. Czik, and E. Bonek, "What makes a good MIMO channel model," in *Vehicular Technology Conference*, vol. 1, pp. 156 – 160, 2005.
- [4] T. S. Pollock, "Correlation modelling in MIMO systems: when can we Kronecker?," in *Proc. 5th Australian Communications Theory Workshop*, pp. 149 – 153, 2004.
- [5] P. A. Bello, "Characterization of randomly time-variant linear channels," *IEEE Trans. Comm. Sys.*, vol. 11, pp. 360 – 393, 1963.
- [6] P. Teal and T. A. Abhayapala, "Spatial correlation in non-isotropic scattering scenarios," in *Proc. IEEE International Conference on Acoustics, Speech and Signal Processing*, vol. III, pp. 2833 – 2866, 2002.
- [7] T. Abhayapala, T. Pollock, and R. Kennedy, "Spatial decomposition of MIMO wireless channels," in *Proc. IEEE Seventh International Symposium on Signal Processing and its Applications, ISSPA 2003*, vol. 1, pp. 309–312, July 2003.
- [8] A. Graham, *Kronecker products and matrix calculus with applications*. Chichester: Ellis Horwood, 1981.
- [9] N. L. Johnson and S. Kotz, "On some generalized Farlie-Gumbel-Morgenstein distribution," *Communications in Statistics*, vol. 4, pp. 415 – 427, 1975.
- [10] T. Betlehem, T. D. Abhayapala, and T. A. Lamahewa, "Space-time MIMO channel modelling using angular power distributions," in *Proc. 7th Australian Communications Theory Workshop*, (to appear), 2006.
- [11] T. J. Kozubowski and K. Podgorski, "A multivariate and asymmetric generalization of Laplace distribution," *Computational Statistics*, vol. 15, no. 4, pp. 531 – 540, 2000.
- [12] G. J. Foschini and M. J. Gans, "On limits of wireless communications in fading environments when using multiple antennas," *Wireless Personal Communications*, vol. 6, pp. 311 – 335, 1998.
- [13] M. T. Ivrlac and J. A. Nossek, "Quantifying diversity and correlation of Rayleigh fading MIMO channels," in *IEEE International Symposium on Signal Processing and Information Technology, ISSPIT'03*, 2003.